

Downward Wage Rigidity in Europe

A New Flexible Parametric Approach and Empirical Results

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Gliederung

- 1 Introduction
- 2 The parametric approach
- 3 Empirical results
- 4 Conclusion

Downward Nominal Wage Rigidity (DNWR)

Some empirical facts

- High unemployment in Europe
- Low inflation
- Inflexible labour markets
- Controversial discussion about extent and relevance of DNWR

Previous work on DNWR I

- DNWR is fundamental for Phillips-Curve discussion
- 'Inflation greases wheels of labor market'
- During the last 10 years several empirical studies
- General finding: existence of DNWR but small effects on unemployment

Previous work on DNWR II

Empirical studies

- Most cited: Kahn (1997, AER)
- Histogram location approach
- Subsequently replicated by:
 - Christofides and Leung (2003)
 - Holden and Wulfsberg (2004)
 - Knoppik and Beissinger (2005)
 - Behr (2006)

Histogram location approach

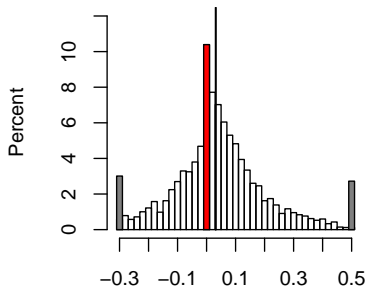
- Assumption that constant share ρ of decreasing wages is constant instead
- Pros
 - no distributional assumption
- Cons
 - assumes identical (only shifted, scaled) distributions over time
 - does not allow for valid inference
 - does not allow to estimate yearly extent of DNWR

Our parametric approach

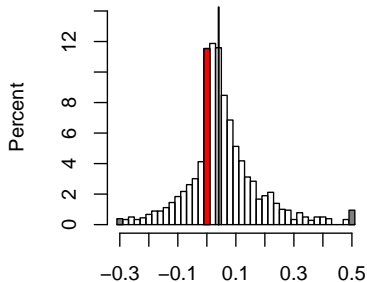
- Assumption that constant share ρ of potential negative wage changes constant due to DNWR
- x is observed wage change, x^C is counterfactual wage change in the absence of rigidity
- Assumption of parametric wage change distribution
- What distribution to choose?

Histograms of wage change distributions 2001, I

Germany

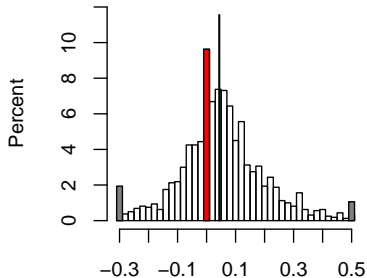


Denmark

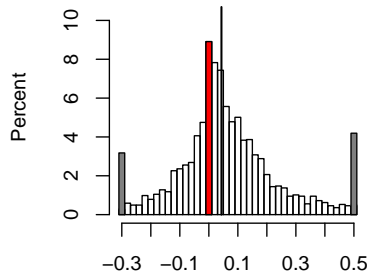


Histograms of wage change distributions 2001, II

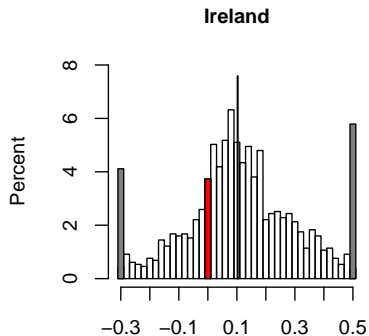
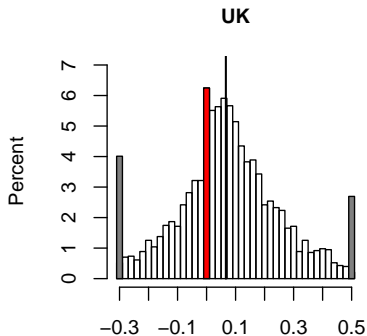
Belgium



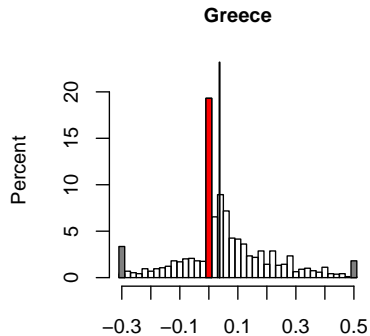
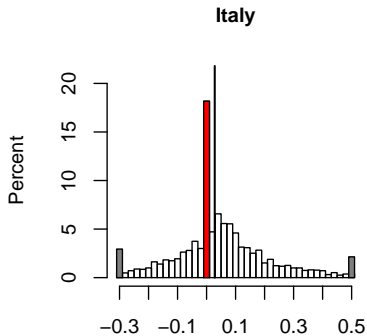
France



Histograms of wage change distributions 2001, III

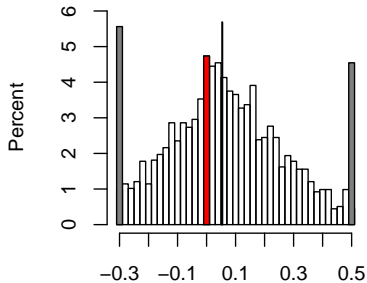


Histograms of wage change distributions 2001, IV

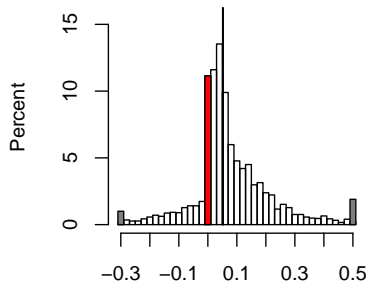


Histograms of wage change distributions 2001, V

Spain



Portugal



Stylized facts of wage change distributions

- unimodal
- asymmetric
- right skewed
- excess kurtosis

Generalized hyperbolic distribution

Characteristics

The generalized hyperbolic distributions

- is specified by five parameters
- has all moments
- can be simulated as Gaussian mixtures using inverse Gaussian as mixture distribution
- can be estimated by ML using numeric methods

Generalized hyperbolic distribution

The density

$$f(x; \Psi) = \kappa \left\{ \delta^2 + (x - \mu)^2 \right\}^{\frac{1}{2}(\lambda - \frac{1}{2})} K_{\lambda - \frac{1}{2}} \left(\alpha \sqrt{\delta^2 + (x - \mu)^2} \right) e^{\beta(x - \mu)}$$

where

$$\kappa = \frac{(\alpha^2 - \beta^2)^{\frac{\lambda}{2}}}{\sqrt{2\pi} \alpha^{\lambda - \frac{1}{2}} K_{\lambda} \left(\delta \sqrt{\alpha^2 - \beta^2} \right)}$$

$K_{\lambda}(\cdot)$ is the modified Bessel function

- α kurtosis parameter
- β skewness parameter
- δ scale parameter
- μ location parameter
- λ Bessel parameter

The rigidity model

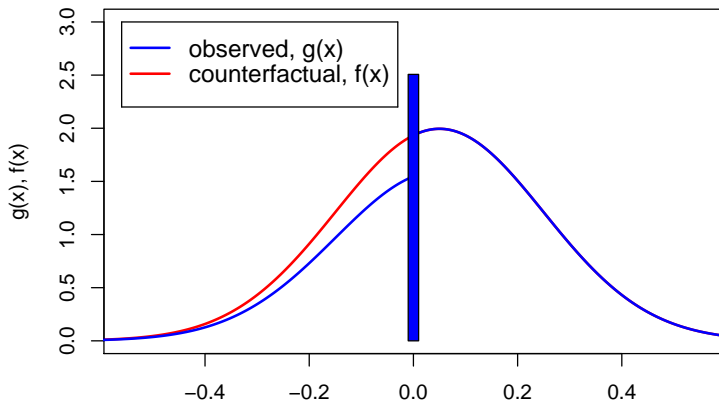
- counterfactual density in the absence of rigidity: $f(x)$
- observed density: $g(x)$
- distributions differ for $\Delta \log(\text{wage}) = x \leq 0$
- Therefore

$$f(x) > g(x) \quad \text{if } x < 0$$

$$f(x) = g(x) \quad \text{if } x > 0$$

Observed and counterfactual distributions

Observed and counterfactual distribution



Relation between $g(x)$ and $f(x)$

$$g(x) = I(x > \epsilon)f(x) + I(x < -\epsilon)(1 - \rho)f(x) \\ + \frac{I(-\epsilon \leq x \leq \epsilon)}{2\epsilon} \left[\int_{-\epsilon}^{\epsilon} f(u) du + \rho \int_{-\infty}^{-\epsilon} f(u) du \right]$$

with indicator function $I(\cdot)$

We estimate the vector of parameters $(\alpha, \beta, \delta, \mu, \lambda)'$ by maximizing numerically the Log Likelihood

$$\log L(\rho, \Psi | x) = \sum_{i=1}^n \log \left(\begin{array}{c} I(x_i > \epsilon) f_{GH}(x_i; \Psi) \\ + I(x_i < -\epsilon) (1 - \rho) f_{GH}(x_i; \Psi) \\ + \frac{I(-\epsilon \leq x_i \leq \epsilon)}{2\epsilon} \\ \left[\int_{-\epsilon}^{\epsilon} f_{GH}(u; \Psi) du + \rho \int_{-\infty}^{-\epsilon} f_{GH}(u; \Psi) du \right] \end{array} \right)$$

Simulation results

- Simulations based on Gaussian mixtures
- Results show that
 - ρ is estimated consistently
 - ρ is estimated with small standard errors
 - standard error of ρ is estimated with high precision

Meaning of ρ

- ρ is conditional relative frequency

$$\rho = \Pr(X = 0 | X^c < 0)$$

- Share of employees potentially 'constraint' is

$$\omega = \Pr(X^c < 0)$$

- Share of 'constraint' employees is

$$\eta = \rho\omega$$

- ω and therefore η varies with rate of inflation π

Overall wage increase due to rigidity

- Rigidity prevents some wages to fall
- Overall wage increase is higher than in the absence of rigidity
- The model allows to estimate the extent of wage increase

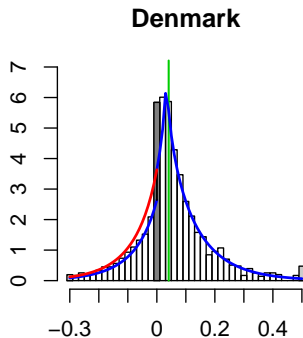
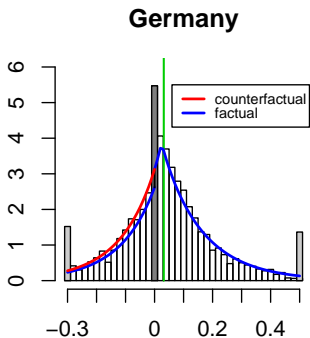
Estimating wage increase due to rigidity

The overall wage increase can be estimated in two ways:

- Based on empirical wages and estimated ρ
- Based on estimated parametric wage distribution and ρ

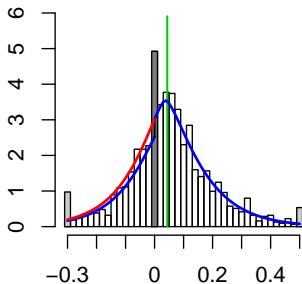
$$E(X^c) = \frac{E(X|X < 0)}{1 - \rho} \Pr(X < 0) + E(X|X > 0) \Pr(X > 0)$$

Observed and counterfactual distributions 2001, I

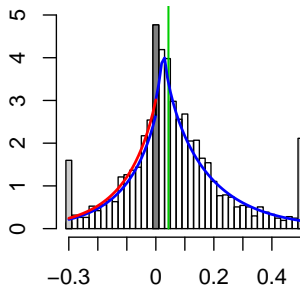


Observed and counterfactual distributions 2001, II

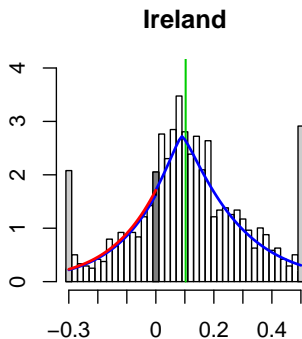
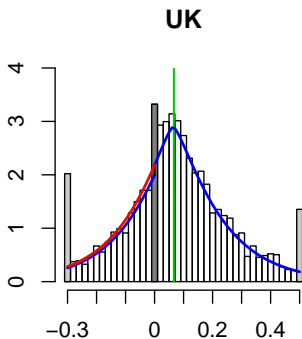
Belgium



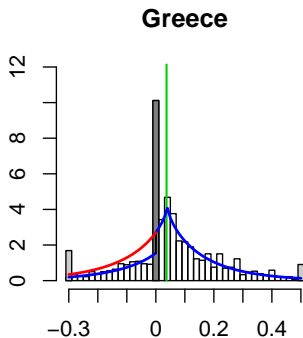
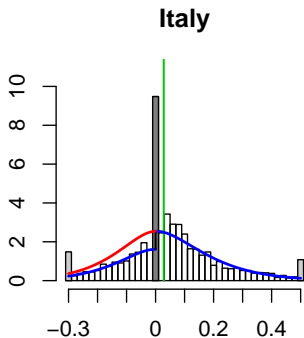
France



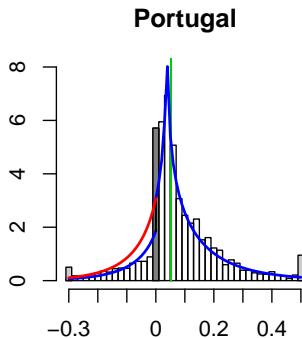
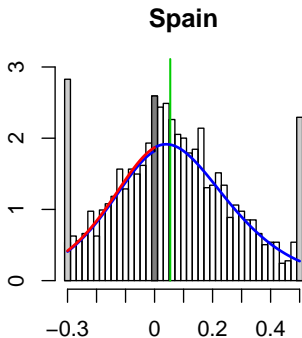
Observed and counterfactual distributions 2001, III)



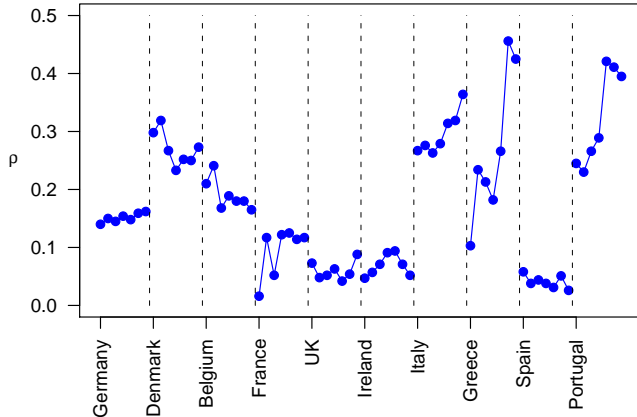
Observed and counterfactual distributions 2001, IV



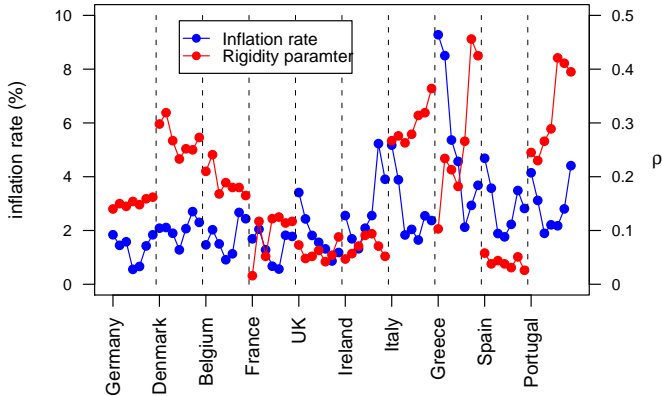
Observed and counterfactual distributions 2001, V



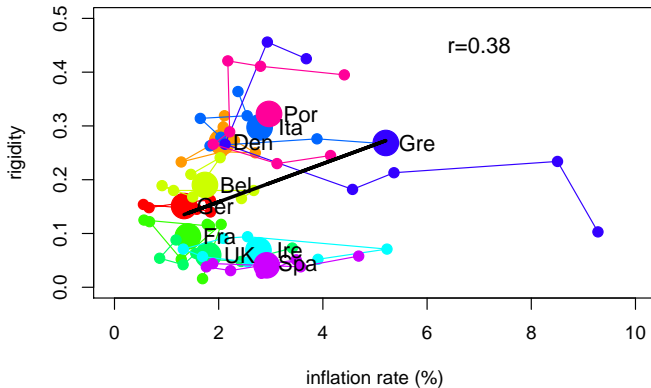
Estimates of rigidity coefficient ρ



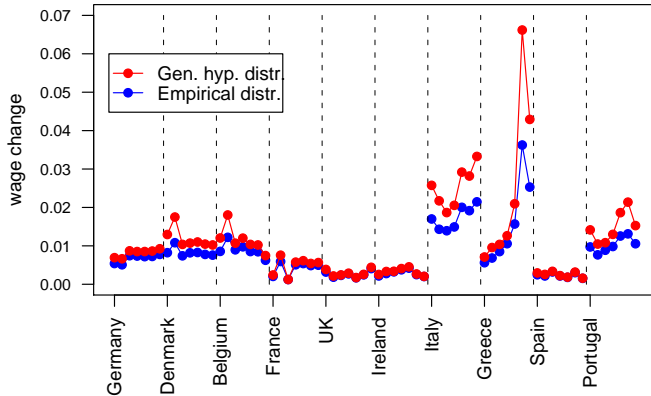
Inflation rates and DNWR I



Inflation rates and DNWR II







Wage increase due to DNWR



Conclusion

- Parametric approach using the Generalized hyperbolic distribution
- allows yearly estimates of rigidity parameter
- strong differences in extent and change of rigidity parameters
- differences are somewhat dampened by positive correlation with inflation rates
- smaller differences in inflation rates will cause higher costs of high rigidity (ρ)

References

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